

## QUESTION ONE (30 MARKS)

- (a) Let  $\pi_i$  be the probability of including unit  $i$  in the sample. Define  $\alpha_i = \{1, \text{ if the } i\text{th unit is in the sample and } 0 \text{ otherwise.}$  Further, let  $Y$  and  $X$  be the survey variable and auxiliary variable respectively.

Prove that the Horvitz-Thomson estimator  $\bar{y}_{HT} = \sum_{i \in s} \frac{y_i}{N\pi_i}$  is unbiased for  $\bar{Y}$  [3 Marks]

- (b) (i) Derive the simplification for the variance of the Horvitz-Thomson  $Var(\bar{y}_{HT})$  due to Yates and Grundy. [3 Marks]  
 (ii) Explain a drawback for the variance estimator due to Yates and Grundy. [2 Marks]

- (c) In a population with  $N = 6$  the values of  $y_i$  are 8, 3, 1, 11, 4 and 7. Calculate the sample mean  $\bar{y}$  for all possible simple random samples of size 2. Verify that  $\bar{y}$  is an unbiased

estimate of  $\bar{Y}$  and show that  $var(\bar{y}) = \left(\frac{N-n}{n}\right) \frac{S^2}{n}$ ;  $S^2 = \frac{\sum_{i=1}^n (y_i - \bar{Y})^2}{N-1}$ . [4 Marks]

- (d) For  $N = 4$ ,  $n = 2$  Consider the estimator  $\hat{Y}_{12} = \frac{1}{2}y_1 + \frac{1}{2}y_2$ ,  $\hat{Y}_{13} = \frac{1}{2}y_1 + \frac{2}{3}y_3$ ,  $\hat{Y}_{23} = \frac{1}{2}y_2 + \frac{1}{3}y_3$ , where,  $\hat{Y}_{ij}$  is the estimator for the sample that has units  $(i, j)$ . Prove that  $\hat{Y}_{ij}$  is unbiased and that  $V(\hat{Y}_{ij}) < V(\bar{y})$  if  $y_3(3y_2 - 3y_1 - y_3) > 0$ . [4 Marks]

- (e) Let  $Y_1, Y_2, \dots, Y_N$  be population values from which a simple Random Sample of size  $n$  is picked under a no Replacement Scheme. Define an estimator of the population mean as  $\hat{Y} = \sum_{i=1}^n \left(\frac{y_i}{N\pi_i}\right)$  where  $\pi_i$  is the inclusion probability of the  $i^{th}$  unit in the said population.

(i) Show that the defined estimator in its form is unbiased for the population mean  $\bar{Y} = \frac{1}{N} \sum_i Y_i$  [5 Marks]

(ii) Derive the variance of the defined estimator and briefly discuss it. [4 Marks]

- (f) Suppose that from certain population  $H$ , there are  $r_h$  respondents with  $n_h$  non-respondents. A sampler wishes to assign a value to every missing one. In this case, explain ny use of appropriate mathematical expressions and or equations, how the sampler would perform;

(i) Random Imputation [2 Marks]

(ii) Nearest Neighbour Imputation [3 Marks]

## QUESTION TWO (20 MARKS)

- (a) Let a population consist of  $N$  units and suppose that the population is divided into  $k$  mutually disjoint strata in such a way that the  $i^{th}$  stratum has  $N_i$  units. A surveyor requires a sample of size  $n$  from this stratified population where  $Y_{ij}$  represents the  $j^{th}$  observation in the  $i^{th}$  stratum. Being that from the  $i^{th}$  stratum, a sample of size  $n_i$  is picked with  $y_{ij}$  being the  $j_{th}$  sampled measurement from the  $i_{th}$  stratum.

(i) Let  $\bar{y}_w = \sum_i^k \left(\frac{N_i}{N} \bar{y}_i\right)$ , Where  $\bar{y}_i$  is the mean of the sampled units from the  $i^{th}$  stratum.

Show that  $\bar{y}_w$  is an unbiased estimator of the population mean  $\bar{Y}$ , derive its variance and discuss the circumstances under which it would vanish. [10 Marks]

- (ii) The sampler during the pilot phase tried sampling from this population and realized that using a simple mean  $\bar{y} = \frac{1}{n} \sum_i^N Y_i$ ,  $\text{var}(\bar{y}) = \left(\frac{1}{n} - \frac{1}{N}\right)S^2$

Whereas in using stratified sampling and assigning sample sizes to every stratum optimally, with a fixed overall sample size  $n$ .

$$\text{var}(\bar{y}) = \frac{1}{n} \left( \sum_i^N \frac{N_i}{N} S^2 \right) - \frac{1}{N} \left( \sum_i^N \frac{N_i}{N} \right) S^2$$

By comparing the two variances, advice the surveyor, with reasons, on the best scheme to adopt. [10 Marks]

### QUESTION THREE (20 MARKS)

- (a) Assume that a population has  $NM$  elements grouped in  $N$  clusters each with size  $M_i$ . Suppose that one picks a sample of size  $n$  from this population and wishes to estimate the population mean, say,  $\bar{Y}$  using this sample. The estimators  $\hat{Y}_1 = \frac{1}{n} \sum_i^n \bar{y}_i$  and  $\hat{Y}_2 = \frac{1}{n} \sum_i^n \frac{M_i}{M} \bar{y}_i$  are proposed, where  $\bar{y}_i$  is the mean of the sampled values from the  $i^{\text{th}}$  cluster. Derive the variances of the two estimators and choose the best one of the two citing reasons for your choice. [20 Marks]

### QUESTION FOUR (20 MARKS)

- (a) Briefly explain with the aid of relevant and logical examples, the concept of;
- (i) Repeated sampling
  - (ii) Double sampling
- as used in studies of finite population [6 Marks]
- (b) Consider a fixed population model in which the population under study is split into two parts, the respondents  $R$  and the non-respondents  $M$  with  $r$  denoting the sample from the respondents and  $m$  representing the sample from the non-respondents. Defining  $y_1, y_2, y_3, \dots, y_r$  as the observations from the respondents with a sample mean of  $\bar{y}_r = \frac{1}{r} \sum_{i=1}^r y_i$  derive an expression for the bias of  $\bar{y}_r$  as an estimator of the population mean  $\bar{Y}$  and briefly discuss your result. [8 Marks]
- (c) One of the main drawbacks of systematic sampling is that it's a Simple Random Sample with only one cluster. Because of this, it is never possible to propose an estimator of the variance based on such a sample. Carefully explain a possible method that can be employed to address this problem and provide a real life example to illustrate your proposition. [6 Marks]